

Uncertainty Quantification in Bayesian Inversion

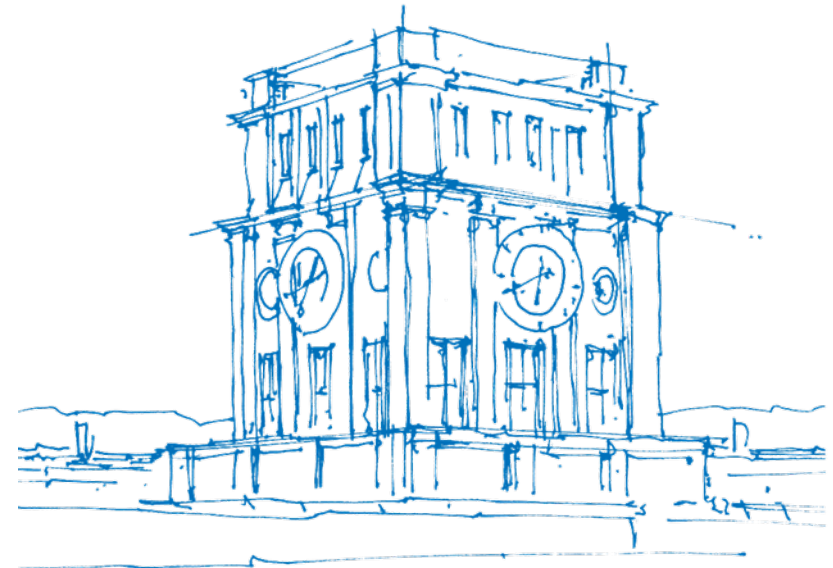
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November 17, 2016



TUM Uhrenturm

Outline

- Brief introduction to Bayesian Statistics and Bayesian Inverse Problems
- Andrew M. Stuart - Uncertainty Quantification in Bayesian Inversion.

Bayesian Statistics

Setting:

- Observation: data point $y \in \mathbb{R}^n$,
- Parameterized distribution of y : $\mathbb{P}(y \in \cdot | u)$, given by a Lebesgue-density

$$\frac{d\mathbb{P}(y \in \cdot | u)}{d\lambda^n}(x) =: L(x|u),$$

- Parameter $u \in X$.

Task: Identify the parameter u based on the observed data point y .

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Posterior: $\mu_{\text{post}} := \mathbb{P}(u \in \cdot | y)$ (knowledge/distribution after seeing the data)

Bayesian Statistics: How to derive the posterior?

Under further assumptions, the posterior can be derived using Bayes' formula:

$$\frac{d\mu_{\text{post}}}{d\mu_{\text{prior}}}(u) = \frac{L(y|u)}{\int L(y|u)d\mu_{\text{prior}}(u)} =: \frac{L(y|u)}{Z(y)}.$$

Assume $u \in X = \mathbb{R}^k$ and μ_{prior} has a pdf f_{prior} , then:

$$f_{\text{post}}(u) = \frac{L(y|u)}{\int L(y|u)d\mu_{\text{prior}}} f_{\text{prior}}(u) = \frac{L(y|u)}{Z(y)} f_{\text{prior}}(u)$$

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Computational: Produce (weighted) samples of the posterior to approximate it empirically.

- Importance Sampling (requires to estimate $Z(y)$)
- Markov Chain Monte Carlo (MCMC; does not yield independent samples of μ_{post})
- Sequential Monte Carlo (SMC; efficient, even if posterior is multimodal or concentrated; requires to estimate $Z(y)$)

Inverse Problem

- $\mathcal{G} : X \rightarrow Y$ is the forward response operator,
- $\eta \sim \mathcal{N}(0, \Gamma)$ is noise,
- $u^{\text{truth}} \in X$ is the true model parameter,
- $y \in Y$ is (noisy) observed data of the model, i.e. given by $y := \mathcal{G}(u^{\text{truth}}) + \eta$.

Identify the parameter u^{truth} , based on the data y .

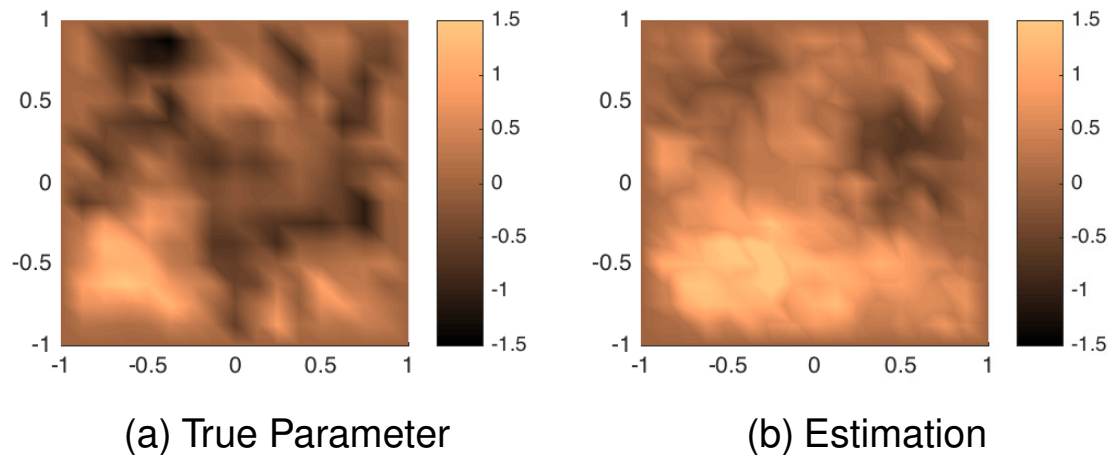


Figure: log-Permeability of an Oil Reservoir

Bayesian Inverse Problem

Let $u \sim \mu_0 := \mu_{\text{prior}}$ and (u, η) independent. Then,

$$\mathcal{G}(u) + \eta = y \Leftrightarrow \eta = y - \mathcal{G}(u),$$

and therefore,

$$y - \mathcal{G}(u) \sim \mathbf{N}(0, \Gamma).$$

The likelihood is then given by

$$L(y|u) := \phi_{0, \Gamma}(y - \mathcal{G}(u)) := \exp\left(-\frac{1}{2} \|\Gamma^{-\frac{1}{2}}(y - \mathcal{G}(u))\|_2^2\right)$$

The posterior $\mu^y := \mu_{\text{post}}$ is then given by Bayes' formula:

$$\mu^y = \frac{1}{Z(y)} \exp\left(-\frac{1}{2} \|\Gamma^{-\frac{1}{2}}(y - \mathcal{G}(\cdot))\|_2^2\right) \mu_0.$$

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